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1. INTRODUCTION
In practical wireless systems, training sequences or pilot symbols are incorporated in the transmitted signal to allow for channel estimation at the receiver. The density of training data needs to increase as the mobility and the channel variation increases. Nevertheless, even with training data available, the channel estimate can only be of limited quality and the channel estimation errors reduce the channel capacity. Furthermore, the fact of substituting data to be transmitted by training data obviously also limits the capacity. All this means that the channel capacity degrades with mobile speed and to minimize this decrease, all a pri-

ori information about the channel should be exploited for its estimation.

In order to exploit partial Channel State Information at the Transmitter (CSIT) in MIMO systems, most of the current precoding schemes exploit either mean [1] or covariance [2] information. A combination of the two can improve exploitation of channel knowledge by weighting them according to a certain criteria. However, mean and covariance information do not necessary have to correspond to actual channel mean and covariance, i.e. to prior distributions. They can be given by a Bayesian approach, with a certain posterior mean and covariance.

In our work, we present different techniques to combine mean and covariance information. We show how to exploit both sources of partial CSIT to optimize the error rate in MIMO systems, by performing linear precoding at the transmitter.

2. (MIMO) WIRELESS CHANNEL MODEL
The impulse response of the time-varying channel is represented by the matrix $h(t)$, which has dimension $N_R \times N_T$, being $N_R$ and $N_T$ the number of receive and transmit antennas respectively. In our work we consider both separable and non-separable models. In a separable channel model the spectrum of the temporal variation in the case of a diffuse channel can then be written as

$$S_{hh}(f) = C_{\tau} \otimes C_T \otimes C_R \cdot S_d(f)$$

where
- $C_{\tau}$: covariance matrix between delays, typically diagonal with power delay profile
- $C_T$: TX side spatial covariance matrix
- $C_R$: RX side spatial covariance matrix
- $S_d(f)$: scalar common Doppler spectrum of all impulse response coefficients. A separable model in 4D is admittedly unrealistic in practice.
In order to simplify the theoretical analysis, a flat MIMO channel will be considered and we shall ignore its Doppler characteristics. In this section, we introduce some different cases of partial CSIT combining mean and covariance information, both for models with separable and non-separable covariance information.

3.1. Non-separable covariance case

When mean and covariance information are present at the transmitter side, this information can be of different nature. In the presence of a Line-of-Sight component between transmitter and receiver, the MIMO channel may be modeled as Ricean. In this first case of interest, the vectorized channel will be considered and we shall ignore its Doppler characteristics. In this section, we introduce some different cases of combining mean and covariance information at the transmitter side. This information can be of different nature: estimation noise, quantization noise and prediction noise. The combination of mean and covariance information leads to a Gaussian posterior distribution with posterior mean given by

\[ \hat{\mathbf{h}} = (\mathbf{I}_{N_R N_T} + \sigma_h^2 \mathbf{C}^{-1})^{-1} \mathbf{h} \]  

and posterior covariance

\[ \hat{\mathbf{C}} = (\sigma_h^{-2} \mathbf{I}_{N_R N_T} + \mathbf{C}^{-1})^{-1} \mathbf{C} \]  

The same noisy mean information can be also combined with noisy covariance information, as a result of quantization noise (limited rate feedback) or estimation noise. Depending on the modeling approach, it could also lead to a Gaussian posterior distribution with a certain mean and covariance. Another case of interest is the noisification of the mean. All previous cases lead to a Gaussian distribution for the channel, be it prior or posterior, with a certain mean \( \hat{\mathbf{h}} \) and covariance \( \hat{\mathbf{C}} \). This information gets reduced by a noisy mean \( \tilde{\mathbf{d}} \) where \( d \) is \( \mathcal{CN}(0, 1) \). This means that the mean becomes zero and the covariance becomes \( \tilde{\mathbf{R}} = \mathbf{hh}^T + \hat{\mathbf{C}} \), which was the correlation matrix before the noisification of the mean. Combined with (3) this gives

\[ \tilde{\mathbf{R}} = \mathbf{hh}^T + (\sigma_h^{-2} \mathbf{I}_{N_R N_T} + \mathbf{C}^{-1})^{-1} \]  

This approach may be of interest since may lead to a simpler problem formulation.

3.2. Separable covariance case

Now consider the channel impulse response matrix \( \mathbf{h} \) with dimension \( N_R \times N_T \) and separable covariance. Hence

\[ \mathbf{E} \mathbf{h} \mathbf{h}^H = \text{tr}(\mathbf{C}) \mathbf{C}_R \]  

In the Ricean case, the channel can be modeled as \( \mathbf{h} = m_{\mathbf{h}} + \mathbf{C}_R^{1/2} \mathbf{h}_w \mathbf{C}_T^{1/2} \) with \( \mathbf{h}_w \) distributed as \( \mathcal{CN}(0, 1) \). A special case can be considered in the Ricean model when also the mean is separable. Thus, the vectorized mean can be represented as \( \tilde{\mathbf{h}} = m_{\mathbf{h}} \otimes m_{\mathbf{R}} \) and matricially as \( \mathbf{m}_{\mathbf{h}} = m_{\mathbf{R}} m_{\mathbf{R}}^H \). In the case of noisy channel estimates and separable perfect covariance information, the posterior mean and covariance are respectively given by

\[ \hat{\mathbf{h}} = (\mathbf{I}_{N_R N_T} + \sigma_h^2 \mathbf{C}_R^{-1} \otimes \mathbf{C}_T^{-1})^{-1} \mathbf{h} \]  

and

\[ \hat{\mathbf{C}} = (\sigma_h^{-2} \mathbf{I}_{N_R N_T} + \mathbf{C}_R^{-1} \otimes \mathbf{C}_T^{-1})^{-1} \]  

If only posterior mean is present, it is due to noise-free channel estimation \( (\sigma_h^2 \to 0) \) and thus \( ||\mathbf{C}|| \to 0 \). On the other hand, only posterior covariance will be present if \( \hat{\mathbf{h}} = \hat{\mathbf{h}} = 0 \) or the estimation noise tends to infinity. In addition, if a rich scattering environment is assumed at the Rx side, the covariance at the receiver can be modeled as identity. In this case, the posterior mean is given by

\[ \hat{\mathbf{h}} = \mathbf{h}(\mathbf{I}_{N_R} + \sigma_h^2 \mathbf{C}_T^{-1})^{-1} \]  

and the posterior covariance

\[ \hat{\mathbf{C}} = \mathbf{I}_{N_R} \otimes \hat{\mathbf{C}}_T \]  

with posterior covariance seen from the transmitter

\[ \hat{\mathbf{C}}_T = (\sigma_h^{-2} \mathbf{I}_{N_T} + \mathbf{C}_T^{-1})^{-1} \]  

In a simplified scenario with noisification of the mean, assume we only have access to \( \mathbf{D} \mathbf{h} \) instead of having access to \( \tilde{\mathbf{h}} \) directly, where the elements of the diagonal matrix \( \mathbf{D} \) are i.i.d. \( \mathcal{CN}(0, 1) \). Now the distribution becomes zero mean with transmit side covariance matrix \( \tilde{\mathbf{R}}_T = \mathbf{h} \mathbf{h}^T + \hat{\mathbf{C}}_T \). Under these circumstances, the mean information falls into the covariance information, and thus the correlation becomes the covariance. Hence, optimal MIMO transmission schemes with partial CSIT for the case of only covariance information will apply for the described model that combines mean and covariance information. If linear prefiltering
is carried out at the transmitter side (after the ST encoding stage) to adapt the transmission to the channel knowledge, an optimal prefilter will lead to capacity maximization or typically Pairwise Error Probability (PEP) minimization. If we assume \( C_R = I_{N_R} \), the optimal prefilter that maximizes capacity and minimizes PEP pours power along the eigenvectors of the posterior correlation matrix seen from the transmitter, following a waterfilling power allocation policy for minimum PEP [2] and possibly different weighting for the capacity maximization solution [6].

### 4. Linear Precoding for Error Rate Minimization

In this section, we derive an optimal precoding strategy for error rate minimization in MIMO systems combining mean and covariance information at the transmitter. The source of mean and covariance information can be either prior or posterior, as described in the previous section. We optimize the performance of the proposed system in terms of PEP averaged over \( h \), prior or posterior, with a certain distribution \( \mathcal{CN}(m_h, C_{hh}) \). In the analysis, we follow the work developed by Jongren et al. in [4]. We assume the separable channel model described in (1) and identity covariance matrix at the receiver side.

The PEP is defined as the error probability of choosing the nearest distinct codeword \( C^i \) instead of \( C^j \). The code error matrix can be defined as \( \hat{E} := [C^i - C^j] \). In practice, the average PEP is limited by the minimum distance code error matrix, given by \( \bar{E} = \arg \min E \{ \det [E(i,j)E^H(i,j)] \} \).

The average PEP is given by

\[
P(C^i \rightarrow C^j) = \int P(C^i \rightarrow C^j | h)p_h(h)dh \tag{11}\]

where the complex Gaussian PDF \( p_h(h) \) is

\[
p_h(h) = \frac{e^{-tr[(h - m_h)C_{hh}^{-1}(h - m_h)]}}{\pi^{N_R/2} \det(C_{hh})^{N_R/2}} \tag{12}\]

By applying the Chernoff bound and averaging over the distribution of \( h \), an upper bound on the average PEP is given by

\[
P(C^i \rightarrow C^j) \leq \frac{1}{2} \int e^{-d_{min}(C^i, C^j)/4}p_h(h)dh \tag{13}\]

When concatenating the Space-Time encoder at the transmitter with a linear prefilter to exploit partial CSIT, the minimum Euclidean distance is

\[
d_{min}^2(C^i, C^j) = d^2(E) = \frac{1}{\sigma^2} \| hWE \|^2 \tag{14}\]

where \( W \) is the linear prefilter. On the other hand, it can be shown that if \( EE^H = \alpha I \), the PEP is minimized at high SNR for a given optimal prefilter [3]. Thus, the system under consideration has \( EE^H = \alpha I \), e.g. orthogonal ST block codes [7] (single stream) or ST spreading [5] (full stream). Introducing \( \eta = \frac{\alpha}{\sigma^2} \) and \( \Psi = WW^H \), the solution to (13) is given by

\[
P(C^i \rightarrow C^j) \leq \frac{\text{tr}[m_h C_{hh}^{-1}(\eta \Psi + C_{hh}^{-1})^{-1} - C_{hh}^{-1}] m_h^H]}{2 \det(\eta \Psi + C_{hh}^{-1})^{N_R} \det(C_{hh})^{N_R}} \tag{15}\]

The performance criterion can be expressed logarithmically (neglecting parameter-independent terms) as follows

\[
J = \text{tr}[m_h C_{hh}^{-1}(\eta \Psi + C_{hh}^{-1})^{-1} - C_{hh}^{-1}] m_h^H - N_R \log \det(\eta \Psi + C_{hh}^{-1}) \tag{16}\]

Assuming a normalized average power constraint, the optimal \( \Psi \) that minimizes the performance criterion in (16) is given by (see Appendix A)

\[
\Psi = \begin{cases}
\frac{1}{2\mu} \left[ 4R[I_{N_T} + \frac{N_R^2 2}{\eta} I_{N_T} + \frac{4}{\eta} m_h C_{hh}^{-1} m_h^H] \right]^{1/2} & \text{if } \eta > \frac{4R}{N_R^2 + 2}
\end{cases} \tag{17}\]

where \( \mu \) is the Lagrange multiplier associated with the power constraint and \( \{ \cdot \}^+ \) takes the positive semidefinite part. It can be seen straightforward from (17) that as \( \eta \) tends to infinity (i.e. SNR tends to infinity), the optimal \( \Psi \) tends to \( \Psi = \frac{1}{\eta} I_{N_T} \), since in this particular case the value of the Lagrange multiplier is \( \mu = N_R I_{N_T} \). This result is equivalent to transmission without CSIT, which shows that as the SNR increases the importance of CSIT gets reduced. Another solution assuming full-rank \( \Psi \) is provided, to have a more intuitive idea of the unequal power-loading policy at the transmitter. Defining the eigenvalue decompositions \( C_{hh}^{-1} = \Sigma_U^H \Lambda U^H \) and \( WW^H = \Sigma_U^H V^H \), \( V \) and \( U \) unitary matrices and \( \Sigma = \text{diag}(\sigma_1, \sigma_2, \ldots, \sigma_{N_T}) \), \( \Lambda = \text{diag}(\lambda_1, \lambda_2, \ldots, \lambda_{N_T}) \), the solution is given by

\[
\Psi = U \Lambda U^H - \frac{1}{\eta} C_{hh}^{-1} \tag{18}\]

where the diagonal elements in the matrix of eigenvalues \( \Lambda \) are given by (see Appendix B)

\[
\lambda_i = \frac{N_R + \sqrt{N_R^2 + 4 \mu \sigma_i^2}}{2\mu} \tag{19}\]

To obtain the optimal precoder either from (17) or (18), let the eigenvalue decomposition of \( \Psi \) be \( \Psi = V \Phi \Lambda \Phi^H \). Since \( \Psi = WW^H \), the optimal precoder is \( W = V \Phi \Lambda_1 \). Thus, the optimal transmission strategy as reflected in the above equations corresponds to transmission along the eigenvectors of a combination of mean and covariance and a waterfilling power allocation policy. The first and second term
in (18) are differently weighted depending on the SNR and the covariance information. In the remaining of this section we introduce some particular cases of special interest.

4.1. Zero mean information

When the mean information is zero, it can be seen from equation (17) that in this case \( \Psi \) becomes

\[
\Psi = \left\{ \frac{N_P}{\mu} I_{N_T} - \frac{1}{\eta} C^{-1}_h \right\}_+	ag{20}
\]

The value of the lagrange multiplier can be analytically expressed as

\[
\mu = \frac{N_R N_T}{1 + \frac{1}{\eta} tr(C^{-1}_h)}	ag{21}
\]

It is clear from (20) and (21) that as the SNR increases the covariance information becomes less important, and \( \Psi \) converges to a scaled identity matrix.

4.2. Unit rank mean

A particular case of interest is the case when the mean information has rank one (e.g. the Ricean case). Since \( m_h \) is unit rank, also \( C^{-1}_h m_h^H m_h C^{-1}_h \) becomes unit rank. The mean \( m_h \) can be represented as a combination of a pair of vectors \( s \) and \( t \), \( m_h = s t^H \). The solution for \( \Psi \) in the case of unit rank mean derived from the full-rank solution in (18) is given by

\[
\Psi = [u_1 U_2] \Lambda [u_1 U_2]^H - \frac{1}{\eta} C^{-1}_h
\]

where \( \Lambda = \text{diag}(1 + \frac{C^{-1}_h}{\eta}, 0, \ldots, 0) \), \( u_1 \) is the eigenvector associated with the only non-zero eigenvalue and \( U_2 \) are arbitrary vectors chosen such that the matrix \([u_1 U_2]\) forms an orthonormal basis. It can be seen from (22) that as the SNR increases the solution approaches to beamforming along a single direction, defined by \( C^{-1}_h m_h^H m_h C^{-1}_h \), which is a combination of mean and covariance information.

4.3. Singular covariance information

When the covariance information is singular, it can be modeled as follows

\[
C_{hh} = [X_h X_\perp] \begin{bmatrix} A & 0 \\ 0 & 0 \end{bmatrix} [X_h X_\perp]^H \tag{23}
\]

where \( \perp \) and \( // \) represent singular and non-singular parts respectively. Let \( [m_{h//} m_{h\perp}] = m_h [X_h X_\perp] \) and \( C_{//} \) be the non-singular part of \( C_{hh} \). The optimization problem in this case becomes

\[
\begin{aligned}
J &= \min_\Psi \text{tr} \left[ m_{h//} C^{-1}_{//} \left( \eta \Psi_{//} + C^{-1}_{//} \right)^{-1} C^{-1}_{//} m_{h//}^H \right] \\
&\quad - N_R \log \det(\eta \Psi_{//} + C^{-1}_{//}) - \eta \text{tr}(m_{h\perp} \Psi_{\perp} m_{h\perp}^H) \\
\text{s.t.} \quad \text{tr}(\Psi_{//} + \Psi_{\perp}) = 1
\end{aligned}
\]

where \( \Psi = [\Psi_{//} \Psi_{\perp}]^T \). The objective function \( J \) can be divided in two optimization problems \( J = J_{//} + J_{\perp} \) minimized separately. The power constraints in both cases have to be adjusted so that \( P_{//} + P_{\perp} = 1 \). Hence, each minimization problem has a different power constraint associated. The optimization problem for the singular part is given by

\[
\begin{aligned}
J_{\perp} &= \min_\Psi \text{tr} \left( m_{h\perp} \Psi_{\perp} m_{h\perp}^H \right) \\
\text{s.t.} \quad \text{tr}(\Psi_{\perp}) = P_{\perp}
\end{aligned}
\]

The solution for \( \Psi_{\perp} \) is derived in Appendix C. The precoding solution corresponds to eigenbeamforming in the direction of the eigenvector of \( m_{h\perp}^H m_{h\perp} \) associated with the largest eigenvalue \( \lambda_{\perp,1} \). The result of the objective function is \( J_{\perp} = -\eta P_{\perp} \lambda_{\perp,1} \). The remaining optimization problem for the non-singular part is given by

\[
\begin{aligned}
J_{//} &= \min_\Psi \text{tr} \left[ m_{h//} C^{-1}_{//} \left( \eta \Psi_{//} + C^{-1}_{//} \right)^{-1} C^{-1}_{//} m_{h//}^H \right] \\
&\quad - N_R \log \det(\eta \Psi_{//} + C^{-1}_{//}) - \eta P_{//} \lambda_{//,1} \\
\text{s.t.} \quad \text{tr}(\Psi_{//}) = P_{//}
\end{aligned}
\]

The solution for this part is equivalent to the general solution with waterfilling shown in (17), but with reduced dimension and power constraint due to singularities. On the other hand, an optimal power split solution exits (\( P_{//}, P_{\perp} \)) under certain circumstances such that \( J(P_{//}) = J_{//}(P_{//}) + J_{\perp}(1 - P_{//}) \) is minimized. If \( J(P_{//}) \) has an absolute minimum \( P_{//,\text{opt}} \), there are three different possibilities. If \( 0 < P_{//,\text{opt}} / P_{\text{min}} < 1 \), the optimal power for the non-singular part is \( P_{//,\text{opt}} = P_{//,\text{opt}} / P_{\text{min}} \) and for the singular part \( P_{\perp,\text{opt}} = 1 - P_{//,\text{opt}} \). The solution is a combination of beamforming (in \( \perp \) part) and waterfilling (in \( // \) part). If \( P_{//,\text{opt}} / P_{\text{min}} \geq 1 \) then \( P_{//,\text{opt}} = 1 \) and \( P_{\perp,\text{opt}} = 0 \), and the solution is given by waterfilling in the non-singular part. Finally, if \( P_{//,\text{opt}} / P_{\text{min}} \leq 0 \) then \( P_{//,\text{opt}} = 0 \) and \( P_{\perp,\text{opt}} = 1 \), and the solution is given by beamforming in \( m_{h\perp}^H m_{h\perp} \).

5. CONCLUSIONS

In this paper, techniques for combining mean and covariance information have been presented. Both sources of information can be either prior (e.g. correlated channel with LOS) or posterior (given by a Bayesian approach). We provide a general precoding solution for PEP minimization when
combining both sources of partial CSIT at the transmitter, and analyze some cases of special interest. The results show how mean and covariance information should be combined in order to exploit the available sources of CSIT.

Appendix A.

The optimization problem described in (16) can be expressed as

\[
\begin{align*}
\min_{\Psi, \mu} & \quad \text{tr} \left[ m_h C_{hh}^{-1} (\eta \Psi + C_{hh}^{-1})^{-1} C_{hh} m_h^H \right] \\
\text{s.t.} \quad & \quad \text{tr}(\Psi) = 1
\end{align*}
\]

(27)

The solution is obtained by means of the Karush-Kuhn-Tucker (KKT) conditions. Defining the Lagrangian as

\[
L(\Psi, \mu) = \text{tr} \left[ m_h C_{hh}^{-1} (\eta \Psi + C_{hh}^{-1})^{-1} C_{hh} m_h^H \right] - N_R \log \det((\eta \Psi + C_{hh}^{-1})) + \mu (\text{tr}(\Psi) - 1)
\]

where \( \mu \) is the Lagrange multiplier associated with the equality constraint. Differentiating \( L(\Psi, \mu) \) w.r.t. \( \Psi \) we get

\[
\mu \Phi - \eta N_R \Phi - \eta C_{hh}^{-1} m_h m_h^H C_{hh}^{-1} = 0
\]

(29)

where the change of variable \( \Phi = \eta \Psi + C_{hh}^{-1} \) has been used for clarity. The solution for \( \Psi \) to the quadratic matrix equation described above is given by

\[
\Psi = \frac{1}{2\mu} \left[ \frac{1}{\eta} N_R I_{N_R} + \left( \frac{1}{\eta} C_{hh}^{-1} m_h m_h^H C_{hh}^{-1} \right)^{-1} \right] \eta C_{hh}^{-1}
\]

(30)

where \( \{ \cdot \}_+ \) takes the positive semidefinite part.

Appendix B.

Introducing the eigenvalue decompositions \( C_{hh}^{-1} m_h^H m_h C_{hh} = U \Sigma U^H \) and \( \Psi + \frac{1}{\eta} C_{hh}^{-1} = V \Lambda V^H \) in equation (27) and eliminating constant terms, the minimization problem becomes

\[
\begin{align*}
\min_{\Psi, \mu} & \quad \text{tr} \left[ \frac{1}{\eta} \Lambda^{-1} V^H U \Sigma U^H V \right] - N_R \log \det(\Lambda) \\
\text{s.t.} \quad & \quad \text{tr}(\Lambda) = \beta
\end{align*}
\]

(31)

where \( \beta = 1 + \frac{1}{\eta} C_{hh}^{-1} \) and the properties \( \text{tr}(AB) = \text{tr}(BA) \) and \( V^H V = I_{N_T} \) have been used. Since the solution we seek assumes \( \Psi \) positive semidefinite, also \( \Lambda - V^H C_{hh}^{-1} V \) is assumed PSD. The optimum \( V \) that minimizes the first term can be chosen as \( V = U \) [4].

Let \( \Sigma = \text{diag}(\sigma_1, \ldots, \sigma_{N_R}) \) and \( \Lambda = \text{diag}(\lambda_1, \ldots, \lambda_{N_T}) \). The Lagrangian in this case is given by

\[
L(\lambda_i, \mu) = \sum_{i=1}^{N_T} \left( \frac{1}{\eta} \frac{\sigma_i}{\lambda_i} - N_R \log \lambda_i \right) + \mu \left( \sum_{i=0}^{N_T} \lambda_i - \beta \right)
\]

(32)

where \( \mu \) is the Lagrange multiplier corresponding to the power constraint. Differentiating \( L(\lambda_i, \mu) \) w.r.t. \( \lambda_i \) we get

\[
\lambda_i = \frac{N_R + \sqrt{N_R^2 + 4 \frac{\mu}{\eta}}}{2 \mu}
\]

(33)

Thus, the solution for \( \Psi \) is given by

\[
\Psi = U \Lambda U^H - \frac{1}{\eta} \frac{\sigma_i}{\lambda_i} C_{hh}^{-1}
\]

(34)

Appendix C.

In order to minimize the objective function in (25) subject to the power constraint, introduce the following eigenvalue decompositions: \( m_h^H m_h = V_{m_h} \Lambda_{m_h,1} V_{m_h}^H \) and \( \Psi = V_{\Psi} \Lambda_{\Psi} V_{\Psi}^H \). By applying the following inequality \( \text{tr}(AB) \leq \sum \lambda_i(A) \lambda_i(B) \), it can be seen that (25) is minimized (the trace is maximized) by setting \( V_{\Psi} = V_{m_h} \). Let \( \Lambda_{m_h} = \text{diag}(\lambda_{m_h,1}, \ldots, \lambda_{m_h,N_T}) \) and \( \Lambda_{\Psi} = \text{diag}(\lambda_{\Psi,1}, \ldots, \lambda_{\Psi,N_T}) \) ordered decreasingly. The optimization problem becomes

\[
\begin{align*}
J_\perp & = \min_{\lambda_{\Psi,i}} -\eta \sum_{i=1}^{N_T} \lambda_{\Psi,i} \lambda_{m_h,i} \\
\text{s.t.} & \quad \sum_{i=1}^{N_T} \lambda_{\Psi,i} \lambda_{m_h,i} = P_\perp
\end{align*}
\]

(35)

Clearly, the function described above is minimized (the summation is maximized) if all the power is transmitted along the strongest eigenvalue, \( \lambda_{m_h,1} \). Hence, the solution is given by choosing \( \lambda_{\Psi,i} = P_\perp \) and \( \lambda_{m_h,i} = 0, i = 2, 3, \ldots, N_T \). With this choice, the value of the objective function becomes \( J_\perp = -\eta P_\perp \lambda_{m_h,1} \).

6. REFERENCES


